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A CHARACTERIZATION OF THE WARING DISTRIBUTION.(U)

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# A CHARACTERIZATION OF THE WARING DISTRIBUTION.

By R M KORWAR

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SUMMARY. In this note the Waring distribution is characterized by the following property: For a positive integer-valued random variable  $X$ ,  $P(X=r) = p_r$ ,  $r=1,2,\dots$ , and with a finite mean  $\mu$  define two new random variables  $Y$  and  $Z$  by

$$P(Y=r) = q_r = \left( \sum_{k=r+1}^{\infty} p_k + ap_r \right) / (\mu + a), \quad r=0,1,2,\dots,$$

$$P(Z=r) = q'_r = (r+b)p_r / (1+b), \quad r=1,2,\dots,$$

where  $a \geq 0$  and  $b$  are constants with  $b-a+1 > 0$ . Then  $Z$  and  $Y$  truncated at 0 have the same distribution if and only if  $X$  has a Waring distribution.

Let  $X$  be a positive integer-valued random variable (r.v.) with

$$P(X=r) = p_r, \quad r=1,2,\dots, \tag{1}$$

Let  $X$  have a finite mean  $\mu$ . One can define two new classes of r.v.'s  $Y$  and  $Z$  by

$$P(Y=r) = q_r = \left( \sum_{k=r+1}^{\infty} p_k + ap_r \right) / (\mu + a), \quad r=0,1,\dots, \tag{2}$$

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(2)

$$P(Z=r) = q_1^a (1+b)P_r / (1+b), \quad r=1,2,\dots, \quad (3)$$

where  $a \geq 0$  and  $b > -1$  are constants. Distributions of the type (2) have been defined and studied in the literature. See Johnson and Kotz (1969, p.261). It is natural to ask: For what distributions  $X, Y$  and  $Z$  are essentially the same r.v.s? Since  $Y$  takes the values  $0, 1, \dots$ , while  $Z$  the values  $1, 2, \dots$ , it is necessary to truncate  $Y$  at 0, and then it turns out that the above property is a unique property of the Waring distribution. For our purpose we define the Waring distribution of Irwin (1963) by

$$P(W=r) = (\lambda - c) c^{[r-1]} / \lambda^{[r]}, \quad r=1,2,\dots, \quad (4)$$

where

$$\lambda - c > 1, \quad c > 0;$$

and

$$c^{[r]} = c(c+1) \dots (c+r-1), \quad r=1,2,\dots, \quad c^{[0]} = 1.$$

We are now ready to characterize the Waring distribution(4):

Theorem: Let  $X$  be a positive integer-valued random variable given by(1) with a finite mean  $\mu$ . Define  $Y$  and  $Z$  by (2) and (3) respectively for some constants  $a \geq 0$ ,  $b$  satisfying  $b-a+1 > 0$ . Then  $Y$  truncated at 0 has the same distribution as  $Z$  if and only if  $X$  has the Waring distribution(4).

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Proof: "Only if" part: Let  $q_r$  and  $q'_r$  be given by (2) and (3) for some  $a \geq 0$  and  $b-a+1 > 0$  with  $q'_r = (1-q_0)^{-1} q_r, r=1,2,\dots$ ,

Then substituting for the  $p$ 's in (2) from (3) we have

$$q_r = \{(u+b)/(u+a)\} \left\{ \sum_{k=r+1}^{\infty} q'_k / (k+b) + a q'_r / (r+b) \right\}, r=1,2,\dots, \\ = (u+a)^{-1}, r=0, \quad (5)$$

From (5) we obtain

$$q_r - q_{r+1} = \left[ q'_{r+1} / (r+1+b) + a \{ q'_r / (r+b) - q'_{r+1} / (r+1+b) \} \right] (u+b)/(u+a) \\ r=1,2,\dots,$$

from which we get the recurring relation

$$q'_{r+1} = \{ (r+b-ax)(r+b+1) \} q'_r / \{ (r+b+1+x-ax)(r+b) \}, \quad (6)$$

where

$$x = (u+b) / \{ (u+a)(1-q_0) \} = (u+b)/(u+a-1) \quad (7)$$

Equation (6) yields

$$q'_{r+1} = \frac{(r+b-ax) \dots (1+b-ax)(r+1+b)}{(r+1+b+x-ax) \dots (2+b+x-ax)(1+b)} q'_1 \\ r=1,2,\dots \quad (8)$$

However, from (5) and the assumed relation  $q'_r = (1-q_0)^{-1} q_r$  it follows that

: 4 :

$$q_1' = q_0 (3+b) / \{(1-q_0)(1+b+x-ax)\} \quad (9)$$

Finally from (3) , (5)-(9) we obtain

$$p_r = (1+b-ax)^{[r-1]} / (1+b+x-ax)^{[r]} , r=1,2,\dots \quad (10)$$

which is Waring (4) with

$$c = (1+b-ax) \text{ and } \lambda = (1+b+x-ax).$$

It is easily seen from (5)-(7) and the assumption  $b-a+1>0$  that  $c = (u-1)(b-a+1)/(u+a-1)>0$ , and since  $x>1$  we also have  $\lambda-c=x>1$ .

"If" part: Let the  $p$ 's be given by (4). Then using the Waring expansion  $(\lambda-c)^{-1} = \sum_{r=0}^{\infty} c^{[r]} / \lambda^{[r+1]}$ , which is valid for  $\lambda>c>0$ , we obtain

$$\sum_{k=r+1}^{\infty} p_k = (c+r-1)p_r / (\lambda-c)$$

from which and the fact that  $\lambda-c/(\lambda-c-1) + 1$  one could easily verify that  $q_r' = (1-q_0)^{-1} q_r$  where  $q_r$  ,  $q_r'$  are given by (2) and (3) with  $a = (b-c+1)/(\lambda-c)$  ,  $b$  arbitrary but  $b-c+1>0$ . Since  $\lambda-c$  is assumed to be greater than 1, we have for the above  $a$  and  $b$  that  $a<b-c+1$  or  $0<c<b-a+1$ . This completes the proof of the theorem.

Note that our theorem covers the characterization of the Yule distribution: In the "Only if" part the Yule distribution corresponds to the case  $a=b=0$  and in the "if" part to  $c=1$ .

There is, in fact, a certain connection to the characterization of the Yule distribution given by Krishnaji (1970). Krishnaji showed that  $X$  has the Yule distribution

$$P(x=r) = (d-1)r! / \{(d+1)(d+2)\dots(d+r)\}, \quad r=1,2,\dots \quad (11)$$

where  $d-1 > 0$  if and only if  $X$  and the greatest integer in  $UX$  truncated at 0 have the same distribution. Here  $U$  is a uniform r.v. on  $(0,1)$  independently distributed of  $X$ . Now if in the "Only if" part of our theorem we take  $a=b=0$ , (5) will reduce to

$$q_r = \sum_{k=r+1}^{\infty} q_k' / k, \quad r=0,1,\dots$$

which is  $P([UZ] = r)$  as has been shown by Krishnaji. Here  $U$  is uniform on  $(0,1)$  distributed independently of  $Z$  (and hence of  $X$ ) and  $[x]$  denotes the greatest integer in  $x$ . Thus Krishnaji's result as applied to  $Z$  and  $[UZ]$  yields that  $Z$  has the Yule distribution (11), and  $X$  in turn has the Yule distribution (4) with  $c=1$ . Of course, our "Only if" part gives the same result.

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